Estimating low-rank covariance matrices from tomographic projections for Cryo-EM classification

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Single particle reconstruction in Cryo-EM Imaging process in cryo-electron microscopy



The cryo-EM problem



The heterogeneity problem

What if the molecule has more than one possible structure?



(Image source: H. Liao and J. Frank, Classification by bootstrapping in single particle methods, *Proceedings of the 2010 IEEE international conference on biomedical imaging*, 2010.)

Experimental data: 70S Ribosome

10000 image dataset (130-by-130), courtesy Joachim Frank (Columbia University)

Class 1

Class 2





Morphing video

Problem formulation

- Let X₁,..., X_n ∈ ℝ^p be a set of i.i.d. discrete random variables representing volumes in a Cryo-EM dataset. The number of possible states is C ≪ p.
- Let M_s : ℝ^p → ℝ^q project a volume along a viewing direction R_s ∈ SO(3) and filter an image by the contrast transfer function h_s.
- Set of projection images given by

$$\mathbf{I}_{s} = M_{s}\mathbf{X}_{s} + \mathbf{E}_{s}, \quad s = 1, \dots, n,$$

where $\mathbf{E}_1, \ldots, \mathbf{E}_n \in \mathbb{R}^q$ is a set of i.i.d. Gaussian white noise variables of variance σ^2 .

► Given realizations *I*₁,..., *I_n* of *I*₁,..., *I_n*, goal is to cluster according to corresponding state of *X*₁,..., *X_n*.

Subspace approach

- If X_s has C possible values, volumes reside in a C − 1-dimensional affine space U centered in E[X_s] and spanned by Var[X_s]¹.
- Basic algorithm:
 - 1. Estimate viewing directions *R_s* for all images assuming single molecule.
 - 2. Find $\mathbb{E}[\mathbf{X}_s]$ and $\operatorname{Var}[\mathbf{X}_s]$; construct **U**.
 - 3. Project I_s onto M_s **U** and use coordinates α_s to cluster.
 - 4. Reconstruct a volume for each cluster using standard techniques.
 - 5. Re-estimate viewing directions and repeat (optional).
- Given realizations I_1, \ldots, I_n , estimate $\mathbb{E}[\mathbf{X}_s]$ and $\operatorname{Var}[\mathbf{X}_s]$.

Principal component analysis from noisy projections

¹Penczek et al. (2009), Liao & Frank (2010)

Least-squares estimators

 $\mathbb{E}[\mathbf{I}_{s}] = M_{s}\mathbb{E}[\mathbf{X}_{s}], \quad \operatorname{Var}[\mathbf{I}_{s}] = M_{s}\operatorname{Var}[\mathbf{X}_{s}]M_{s}^{H} + \sigma^{2}I_{q}.$

• Given a set of realizations I_1, \ldots, I_n , define

$$\mu_n = \arg\min_{\mu} \frac{1}{n} \sum_{s=1}^n \|I_s - M_s \mu\|^2 + \lambda \|\mu\|^2,$$

and

$$\Sigma_n = \underset{\Sigma}{\operatorname{arg\,min}} \frac{1}{n} \sum_{s=1}^n \| (I_s - M_s \mu_n) (I_s - M_s \mu_n)^H - (M_s \Sigma M_s^H + \sigma^2 \mathbf{I}_q) \|_F^2 + \lambda \| \Sigma \|^2.$$

For uniform distribution of viewing directions, h_s(ω) = 1, and λ = 0, Σ_n ^{a.s.}→ Var[X_s] as n→∞.

Normal equations

Mean estimator satisfies $A_n\mu_n = b_n$, where

$$A_n = \frac{1}{n} \sum_{s=1}^n M_s^H M_s + \lambda I_q, \text{ and } b_n = \frac{1}{n} \sum_{s=1}^n M_s^H I_s.$$

For covariance, $L_n(\Sigma_n) = B_n$, where $L_n : \mathbb{R}^{p \times p} \to \mathbb{R}^{p \times p}$ is defined by

$$L_n(\Sigma) = \frac{1}{n} \sum_{s=1}^n M_s^H M_s \Sigma M_s^H M_s + \lambda \Sigma$$

and

$$B_n = \frac{1}{n} \sum_{s=1}^n M_s^H (I_s - M_s \mu_n) (I_s - M_s \mu_n)^H M_s$$
$$- \sigma^2 \frac{1}{n} \sum_{s=1}^n M_s^H M_s .$$

Inversion of L_n

- Direct inversion of matrix intractable; costs $O(p^6) = O(N^{18})$, since $p = O(N^3)$ for resolution *N*.
- Katsevich et al (2015) designed a basis in which L_n is approximated by a block-diagonal operator. This requires uniform distribution of viewing directions and no CTF.
- Another approach is to use iterative methods to invert L_n, such as the algebraic reconstruction technique (ART), also known as the Kaczmarz method².
- We use conjugate gradient (CG) method to calculate $L_n^{-1}(B_n)$ through repeated application of L_n .
- ► Non-uniform fast Fourier transforms allow us to calculate L_n in O(nN⁶ log N), so T iterations cost O(nTN⁶ log N).

²Liao et al. (2014)

Convolution operator

The operator M^H_s M_s projects, then backprojects, so it can be written as a convolution. As a result

$$\Sigma \mapsto M_s^H M_s \Sigma M_s^H M_s,$$

is a convolution along the rows and columns of Σ .

Plugging this into the expression for L_n, we have

$$L_n(\Sigma) = \frac{1}{n} \sum_{s=1}^n M_s^H M_s \Sigma M_s^H M_s + \lambda \Sigma$$
$$= \Sigma * \text{Ker} + \lambda \Sigma$$

where Ker is a convolution kernel in six dimensions.

Convolution operator (cont.)

- Precalculating Ker in one pass through the dataset takes O(nN⁶ log N), but once calculated, applying L_n amounts to a convolution costing only O(N⁶ log N).
- Solving $L_n(\Sigma_n) = B_n$ then has computational complexity $O((n + T)N^6 \log N)$.
- We can also approximate convolution by circular convolution. In the Fourier domain, we then solve

$$\widehat{\operatorname{Ker}}\circ\widehat{\Sigma}_n=\widehat{B}_n,$$

where \circ is the elementwise matrix product. Complexity for this method is just $O(nN^6 \log N)$.

Low-rank approximation

- However, these methods all have complexity of at least O(N⁶ log N), severely limiting resolution.
- Katsevich et al (2015) showed that

$$\lim_{n\to\infty}\widehat{\operatorname{Ker}}(\vec{\xi_1},\vec{\xi_2}) = \frac{2}{\|\vec{\xi_1}\times\vec{\xi_2}\|},$$

for $h_s(\omega) = 1$ and uniformly distributed R_s .

- The eigenvalues decay as 1/k, so Ker is approximately low-rank.
- For a small C, Var[X_s] has low rank C − 1. Therefore B_n is also of low rank, since B_n is the elementwise product of two low-rank matrices Var[X_s] and Ker.

Low-rank approximation (cont.)

- Being the sum of rank-one matrices, B_n can be applied to vectors fast.
- Probablistic matrix decompositions techniques thus allow for efficient low-rank approximation³.
- Asymptotically, $1/\widehat{\text{Ker}}(\vec{\xi_1}, \vec{\xi_2}) \approx \frac{1}{2} \|\vec{\xi_1} \times \vec{\xi_2}\|$, so we expect

$$\widehat{\mathrm{iKer}}(\vec{\xi_1},\vec{\xi_2}) = \frac{1}{\widehat{\mathrm{Ker}}(\vec{\xi_1},\vec{\xi_2}) + \lambda}$$

to be smooth and therefore low-rank. The Nyström method provides a low-rank approximation of \widehat{iKer} by calculating a subset of its columns and interpolating⁴.

⁴Halko et al (2011) ⁴Williams & Seeger (2001)

Low-rank approximation (cont.)

We can now calculate

$$\widehat{\Sigma}_n = \widehat{\mathrm{iKer}} \circ \widehat{B}_n,$$

where $\widehat{\Sigma}_n$ is a sum of rank-one matrices, since the elementwise product of two rank-one matrices is also rank-one.

- Using this representation, we can apply standard eigendecomposition algorithms to Σ_n to capture its principal eigenvectors.
- Letting r be the rank of B_n and iKer, complexity is now O(rnN³ log N). This is a significant improvement over previous methods.

Conjugate gradient deconvolution: 70S ribosome

Dataset of 10000 images (130-by-130), 2 classes, courtesy Joachim Frank (Columbia University). Downsampled to N = 17.



Took 18 min on a 2.9 GHz, 16-core CPU with 96 GB of memory.

Clustering accuracy 88% with respect to dataset labeling.

Low-rank deconvolution: 70S ribosome

Dataset of 10000 images (130-by-130), 2 classes, courtesy Joachim Frank (Columbia University). Downsampled to N = 17.



Took 5 min on a 2.9 GHz, 16-core CPU with 96 GB of memory.

Clustering accuracy 84% with respect to dataset labeling.

Conclusions

 Principal component analysis of volumes from noisy projections using a least-squares estimator provides a powerful measure of variability in Cryo-EM volumes.

 Precalculating convolutional kernel coupled with conjugate gradient allows for efficient calculating of least-squares covariance estimator.

 Leveraging the low rank of the covariance matrix and convolution kernel significantly reduces complexity in time and memory.

Bibliography

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